

Deterministic or stochastic seasonality in daily electricity prices?

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Abstract The paper presents an analysis of the seasonality of Italian daily electricity prices. Since the correct detection of the nature, stochastic or probabilistic, of the seasonality is crucial in ARIMA modeling, a test that allows such detection is presented. The application of this test to the Italian daily prices in the years 2008-11 has pointed to the presence of deterministic seasonality in the short run. Nevertheless, seasonality has decreased in the last two years as consequence of a more balanced consumption of electricity over the week.

Key words: Electricity prices, seasonal unit roots, HEGY test

1 Introduction

As is well known, daily electricity prices are characterized by clear seasonal patterns associated with time intervals, such as the day of week or the month. Every attempt to construct a model of these prices has to take this evidence into account. In particular, the correct application of ARIMA models requires a deep analysis and identification of the nature, stochastic or deterministic, of the seasonality present in the prices. The treatment of seasonality in the ARIMA framework is conceptually similar to the treatment of trends: like these, seasonality entails the non-stationarity of the process, and its non-stationary effect has to be removed before modeling the process. More specifically, if the daily effects are constant at corresponding days (every Sunday, every Monday,...), the seasonality can be shaped by a periodic linear function $s(t)$ (*deterministic seasonality*). In this case, the correct treatment consists in extracting the seasonality by regression, and then modeling the non-seasonal prices using an ARIMA model:

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$$\phi(B)[p_t - s(t)] = \theta(B)\varepsilon_t \quad (1)$$

where the autoregressive part of the model $\phi(B)$ can include unit roots characterizing the model as integrated. As a whole, model 1 is also named the *Seasonal Reg-ARIMA* model. On the other hand, if the daily effects are characterized by stochastic variability (*stochastic seasonality*), the correct treatment consists in applying the weekly difference to the prices $\Delta_7 p_t = p_t - p_{t-7}$, and then modeling the differences by an ARIMA model:

$$\phi(B)\Delta_7 p_t = \theta(B)\varepsilon_t \quad (2)$$

In this case, the process is named *Seasonally Integrated*. The two treatments are not interchangeable. In fact, in the case of deterministic seasonality, the seasonal difference is not efficient because it introduces seasonal unit roots into the moving average part of the ARIMA model; in the case of stochastic seasonality, the first treatment does not assure the stationarity in the second moment of the data.

The following section illustrates a test that allows us to detect the kind of seasonality affecting the daily electricity prices; Section 3 reports the application of this test to the Italian daily prices of electricity in the period 2008-11. An analysis of the daily effects (seasonality) will also be provided using suitable ARIMA models.

2 HEGY test

A very common methodology to test for non-stationarity due to seasonality is the procedure developed by Hylleberg, Engle, Granger, and Yoo [1], referred to as the HEGY test. This test was originally derived for quarterly seasonality, but it was also extended for weekly seasonality in data collected on a daily basis by Rubia [2]. Under the null hypothesis, the HEGY test assumes that the relevant variable is seasonally integrated. That means, in the case of daily electricity prices (p_t), that the weekly difference $p_t - p_{t-7} = \Delta_7 p_t$ is assumed to be a stationary process. Since $\Delta_7 = (1 - B^7) = (1 - B)(1 + B + B^2 + \dots + B^6)$, the null hypothesis entails the presence of a single unit root at the zero frequency and three pairs of complex roots at the seasonal frequencies $2\pi k/7$, where $k = 1, 2, 3$ represents the number of cycles per week of each frequency.

According to this fact, the test is based on the following auxiliary regression:

$$\Delta_7 p_t = \alpha + \sum_{d=2}^7 \delta_d D_{d,t} + \sum_{r=1}^7 \alpha_r z_{r,t-1} + \sum_{j=1}^p \phi_j \Delta_7 p_{t-j} + \varepsilon_t^1 \quad (3)$$

where $D_{d,t}$ is a zero/one dummy variable corresponding to the d -th day of the week, and the regressors $z_{r,t}$ are defined as follows:

¹ This is a standard version of the HEGY test for daily data, but it can be extended including trends. Nevertheless, in this case, there is not cause for doing it.

$$\begin{aligned}
z_{1,t} &= \sum_{j=1}^7 \cos(j0) p_{t-j} \\
z_{2k,t} &= \sum_{j=1}^7 \cos(j2\pi k/7) p_{t-j} \text{ for } k = 1, 2, 3 \\
z_{2k+1,t} &= \sum_{j=1}^7 \sin(j2\pi k/7) p_{t-j}
\end{aligned} \tag{4}$$

Each regressor is an orthogonal variable that include only one root of the seven roots included in p_t . For example, $z_{1,t}$ includes only the unit root corresponding to zero frequency (random walk), but not the seasonal roots; $z_{2,t}$ and $z_{3,t}$ include only the seasonal roots corresponding to the frequency $2\pi/7$. Therefore, as in the augmented unit root test of Dickey and Fuller (ADF) [3], the process includes a unit root if the null hypothesis $\alpha_1 = 0$ is accepted against the alternative hypothesis $\alpha_1 < 0$ on the basis of a non-standard t -statistic. Moreover, the test allows us to detect the presence of seasonal roots, i.e. stochastic seasonality. Indeed, the presence of a couple of complex roots corresponding to a seasonal frequency $2\pi k/7$ implies that both the parameters of the regressors associated with that seasonal frequency are zero: $\alpha_{2k} = \alpha_{2k+1} = 0$; $k = 1, 2, 3$. This assumption can be tested by a joint F -test; the distributions of each statistic F_k are not standard, but the critical values are reported in [4]. Finally, the auxiliary regression includes a number of lags of the dependent variable in order to avoid serial correlation in the error term ε_t , that reduces the test size.

3 Analysis of the Italian daily electricity prices

The HEGY test was performed on the 2008-2011 Italian daily PUN² (more specifically the log-PUN). As observable in Table 1, none of the null hypotheses (H_0) are significant at 1% level, i.e. the test point to neither a random walk nor seasonal roots. Nevertheless, the absence of a random walk is not confirmed by the ADF test on the same data. That might mean that the prices process is nearly a random walk, but also the process is not homogeneous over the whole period. Indeed, after performing the HEGY test on the sub-periods 2008-09 and 2010-11, we can note that the statistic t concerning the presence of a random walk gives different signals: the 2008-09 daily prices seem to include a random walk movement, whereas the 2010-11 daily prices do not. Such deductions were confirmed by performing the ADF test on the data (Table 1). The absence of mean-reversion in the first period is a particular case and should be related to the high variation of the oil prices in the same period. On the other hand, the seasonality remains non stochastic in both periods (absence of seasonal roots). According to these findings, the 2008-09 daily electricity prices were

² The PUN is the National Single Price in the Italian electricity market (IPEX). The PUN series are downloadable from the web-site of the Energy Markets Manager: <http://www.mercatoelettrico.org>

Table 1 HEGY and ADF tests

H_0	stat.	2008-11 (sign.)	2008-09 (sig.)	2010-11 (sign.)
$\alpha_1 = 0$	t	-3,739 ***	-1,572	-3,742 ***
$\alpha_2 = \alpha_3 = 0$	F_1	163,127 ***	65,480 ***	84,874 ***
$\alpha_4 = \alpha_5 = 0$	F_2	175,639 ***	66,215 ***	90,891 ***
$\alpha_6 = \alpha_7 = 0$	F_3	232,567 ***	103,114 ***	93,018 ***
ADF test	τ	-2,288	-1,167	-3,371 **

shaped by a Seasonal Reg-ARIMA model, whereas the 2010-11 prices were shaped by a Seasonal Reg-ARMA model.

For reasons of space, Table 2 reports only the daily coefficients of the seasonal regression function $s(t)$. Since the analyzed data are log-prices, each daily coefficient indicates the average per-cent difference between the correspondent daily price and the Sunday price (daily effect). We can note that in period 2010-11 such differences decreased, i.e. the daily effects decreased, and this is due to a more balanced consumption of electricity over the week. Such behavior should be related to the arrival of electricity supply contracts (e.g. Enel Bioraria) that make the consumption of electricity at the week-end more economical than on work days.

Table 2 Seasonal effects

day effect	2008-09 (sign.)	2010-11 (sign.)
Monday	0,149 ***	0,076 ***
Tuesday	0,172 ***	0,093 ***
Wednesday	0,190 ***	0,091 ***
Thursday	0,168 ***	0,097 ***
Friday	0,150 ***	0,080 ***
Saturday	0,103 ***	0,072 ***

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